# Enlisting Workers in Monitoring Firms: Payroll Tax Compliance in Mexico

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  - ► Lowest tax revenue/GDP share in the OECD: 15-20% over study period.
  - ▶ Informal economy estimated at 40+% of GDP (Schneider and Enste, 2000).
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  - Mexican social security agency (IMSS) supposed to cover all private-sector workers; in fact covers 53%.
- Non-compliance of firms with tax regulations is a key element of general weakness of state capacity in many developing countries.

- One well-appreciated dimension of non-compliance: failure to register.
  - Generates a variety of distortions: limited access to credit, limits on employment growth (Gordon and Li, 2009; Levy, 2008).
  - Several governments have implemented policies to reduce registration costs, induce firms to register (Fajnzylber et al., 2011; Bruhn, 2011; Kaplan et al., forthcoming).
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  - Recent papers have examined effect of formalization on firm-level outcomes (McKenzie and Sakho, 2010; de Mel, McKenzie and Woodruff, 2012)
- ▶ In this paper, we focus on an arguably under-appreciated form of non-compliance: under-reporting of wages by registered firms, to evade payroll taxes.

- Why under-appreciated?
  - ► Third-party reporting has been found to be quite accurate in developed countries (Saez, 2010; Kleven, Knudsen, Kreiner, Pedersen and Saez, 2011).
  - ▶ Difficult to study. Data requirements formidable.

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#### Strategy:

- Compare (male) wage distribution in IMSS administrative records to wage distribution for similar workers in household survey who report receiving IMSS coverage.
- Use 1997 pension reform as source of exogenous variation in incentive of employees to ensure accurate reporting by their employers.

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- Disclaimer:
  - ► This is not an endorsement of personal retirement accounts *per se*. Incentives to monitor could be given in a traditional pension system as well.

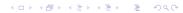
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- ► For time reasons I am not discussing related literature, including a motivating theory model by Kleven, Kreiner and Saez (2009). Please see paper.

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  - Health care: free to covered employees and their families in IMSS clinics and hospitals.
  - Child care: free for children ages 7 weeks-4 years to mothers and single fathers covered in their jobs.
  - Retirement pension (more below)
  - Disability
  - Worker's compensation
  - Housing fund (more below)

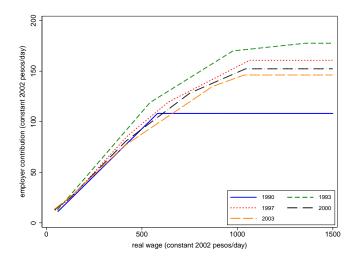
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- Health care, child care, disability, worker's compensation are available to all covered workers, spouses and dependents, independent of wage reported.
- ► Health care, child care, disability, worker's compensation changed little over study period.

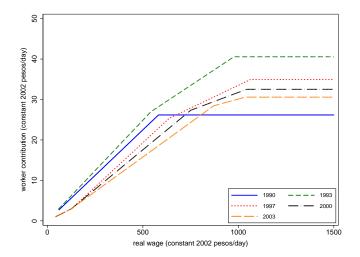


- Contributions:
  - ▶ Employers: 18-23% of wage, for most workers.
  - ► Employees: 2-5% of wage, for most workers. [See figures.]
- Observations:
  - ▶ Changes over time relatively modest.
  - ► Changes affect all age groups similarly; should be differenced out in D-in-D estimation.

### Fig. 1: Employer contribution schedule



### Fig. 2: Employee contribution schedule



#### Pension reform:

- Beginning in 1944, operated as pay-as-you-go (PAYGO) system.
- ightharpoonup Rising number of retirees, macro crises in 1980s ightarrow "fiscal imbalances".
- ▶ In 1992, personal accounts created in parallel with PAYGO system. Plagued by administrative problems.
- ▶ In Dec. 1995, law passed creating new system of personal retirement accounts (PRAs). Implemented July 1, 1997.

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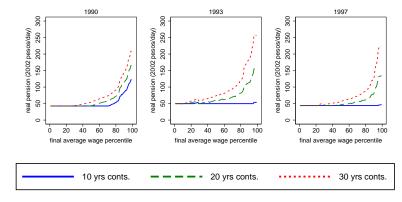
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    - ▶ Gradually raised to 100% of minimum wage in 1995.
  - Many retirees near minimum 10 years of contributions.
  - ▶ Upshot: 80+% of retirees were getting minimum pension prior to 1997 reform. [See figure.]

### Fig. 3: Value of pension, men ages 60-65

C. Value of pension by ENEU wage percentile, ages 60-65



→ Women

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- ► Grandfathering: anyone in the system as of July 1, 1997 retained right to choose (at date of retirement) pension he/she would have had under pre-reform system, calculated as if he/she had always been under the pre-reform system.

#### Fig. 4: Estado de Cuenta



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#### Table 1: Pension wealth simulation

		Real Daily Wage					
Years of	D.	40	100	200	202	500	1070
Contributions	Plan	43	100	200	300	500	1079
35	PRA	399.0	815.0	1626.2	2437.3	4059.7	8759.2
	PAYGO	399.0	672.2	1263.6	1862.8	3104.6	6702.0
30	PRA	399.0	523.4	1044.3	1565.3	2607.1	5625.1
	PAYGO	399.0	594.1	1068.7	1555.4	2592.4	5596.3
25	PRA	399.0	399.0	659.1	987.8	1645.3	3549.9
	PAYGO	399.0	507.0	851.4	1212.7	2021.1	4363.0
20	PRA	88.0	202.4	403.9	605.4	1008.4	2175.7
	PAYGO	399.0	437.9	679.0	940.8	1568.0	3384.8
15	PRA	51.2	117.8	235.0	352.2	586.6	1265.7
	PAYGO	399.0	399.0	484.2	633.5	1055.8	2279.1
10	PRA	26.8	61.7	123.1	184.5	307.4	663.2
	PAYGO	399.0	399.0	399.0	399.0	543.6	1173.4
5	PRA	10.7	24.6	49.0	73.5	122.4	264.2
	PAYGO	0.0	0.0	0.0	0.0	0.0	0.0

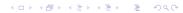
Notes: Values are real present discounted value of the future stream of pension benefits in thousands of 2002 pesos, for a male worker who enters the system on June 30, 1997.

#### Data

- IMSS administrative records:
  - ▶ Full set of employers' reports of employees' wages, 1985-2005.
  - Variables: age, sex, daily wage, state and year of first registration with IMSS, employer id (location, industry)
  - Wages reported as spells; we draw last day of quarter.
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- Encuesta Nacional de Empleo Urbano (ENEU)
  - CPS-like household survey.
  - ► Households survey quarterly for 5 quarters, same questionnaire each period.
  - ▶ Began in 1987, some weirdness in first year.
  - ▶ Initial sample from 16 cities, expanded over time.
  - Questionnaire modified in 1994.
  - ▶ More extensive re-design in 2003.
  - ▶ Asks if workers receive IMSS coverage.
  - Contract type available 1994 on.



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▶ Years: 1988-2003

► Ages: 16-65

Cities: 16 cities in original ENEU sample

- Sectors: manufacturing, construction, retail/hotel/restaurant (sectors in which IMSS is only social security agency.)
- Main (highest-wage) job, if more than one.
- ▶ Impose 1991 IMSS topcode (lowest real value).

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  - Small N problem in ENEU, especially for older women by metro area.
- Summary: cross-sectional results for women similar to those for men. D-in-D noisier, no clear pattern.

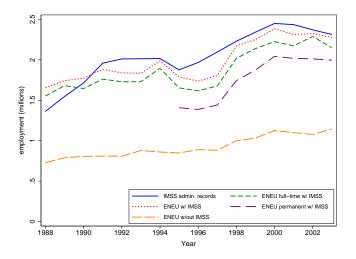


## Table 2: Comparison of IMSS and ENEU, 1990, men

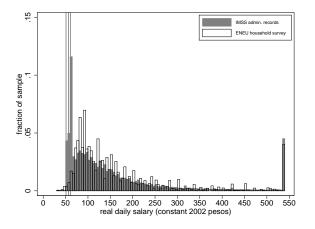
A. 1990	IMSS baseline sample (1)	full ENEU sample (2)	ENEU w/ IMSS (3)	ENEU w/o IMSS (4)	ENEU permanent w/ IMSS (5)	ENEU full-time w/ IMSS (6)
	156.41	16401	170.10	144.00		166.07
real avg daily wage	156.41	164.01	173.12	144.00		166.87
	(0.09)	(1.58)	(1.94)	(2.62)		(1.85)
age	31.81	31.46	32.13	29.98		32.22
	(0.01)	(0.15)	(0.17)	(0.29)		(0.17)
fraction employed in ests >100 employees	0.52	0.43	0.55	0.18		0.55
	(0.00)	(0.01)	(0.01)	(0.01)		(0.01)
N (raw observations)	1714518	16169	11592	4577		10978
N (population, using weights)	1714518	2578847	1772523	806324		1645229
B. 2000						
real avg daily wage	160.28	148.32	161.28	120.88	166.56	155.93
0 , 0	(0.09)	(1.31)	(1.60)	(2.16)	(1.80)	(1.59)
age	32.77	32.22	32.82	30.94	33.22	32.88
_	(0.01)	(0.14)	(0.16)	(0.28)	(0.17)	(0.16)
fraction employed in ests >100 employees	0.58	0.44	0.59	0.10	0.63	0.59
, ,	(0.00)	(0.01)	(0.01)	(0.01)	(0.01)	(0.01)
N (raw observations)	2449442	19171	14063	5108	11918	13246
N (population, using weights)	2449442	3509828	2384267	1125561	2042988	2225318
(population, using weights)	2773442	3309020	2504201	1123301	2072900	2223310



## Fig. 5: Employment, IMSS vs. ENEU samples, men



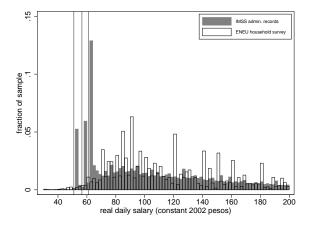
#### Fig. 6: Wage histogram, men, 1990



Notes: Bins are 5 pesos wide. Average 2002 exchange rate: 9.66 pesos/dollar. Vertical lines represent the three region-specific minimum wages.



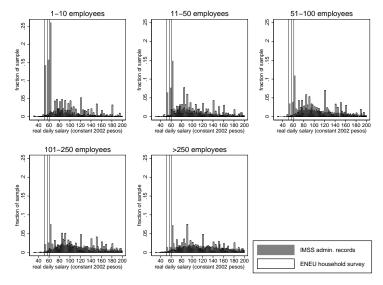
#### Fig. 7: Wage histogram, men, 1990, low wages



Notes: Bins are 2 pesos wide. Average 2002 exchange rate: 9.66 pesos/dollar. Vertical lines represent the three region-specific minimum wages.



## Fig. 8: Wage histograms, men, 1990, by firm size



Notes: Bins are 2 pesos wide. Average 2002 exchange rate: 9.66 pesos/dollar.











Fig. 11: Wage densities by age group, men

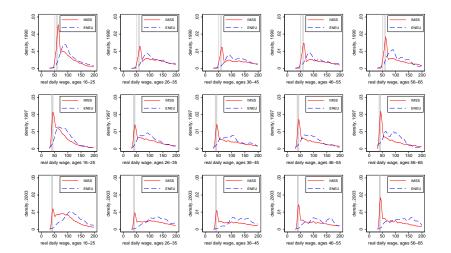


Fig. 12: Wage gaps by age group, men

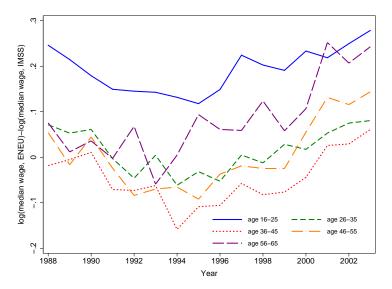
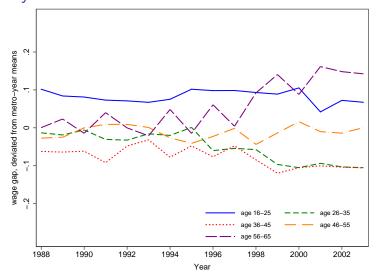




Fig. 13: Wage gaps by age group, men, deviated from metro-year means



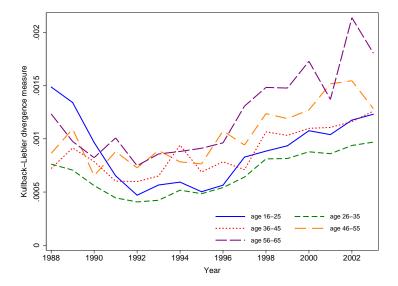
## Kullback-Liebler divergence

▶ Kullback-Liebler divergence between two densities,  $f_1$  and  $f_2$ :

$$D_{12} = \int_0^\infty \frac{f_1(w) - f_2(w)}{f_2(w)} f_1(w) dw$$

- ▶ Only defined for  $f_2(w) > 0$ ; use non-parametric density estimates.
- ▶ Used by DiNardo, Fortin and Lemieux (1996) among others.

Fig. 14: Kullback-Liebler divergence by age group, men

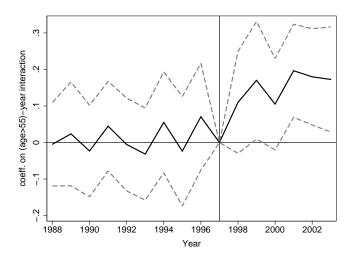




## Table 3: Differential effects on wage gap, men

	dep. var.: log(median wage, ENEU) - log(median wage, IMSS)					
	(1)	(2)	(3)			
1(age > 55)*1988	-0.005	-0.005	-0.005			
, -	(0.088)	(0.065)	(0.058)			
1(age > 55)*1989	0.024	0.024	0.024			
	(0.103)	(0.083)	(0.072)			
1(age > 55)*1990	-0.023	-0.023	-0.023			
	(0.099)	(0.071)	(0.064)			
1(age > 55)*1991	0.045	0.045	0.045			
	(0.106)	(0.070)	(0.063)			
1(age > 55)*1992	-0.005	-0.005	-0.005			
	(0.090)	(0.068)	(0.065)			
1(age > 55)*1993	-0.032	-0.032	-0.032			
	(0.100)	(0.071)	(0.064)			
1(age > 55)*1994	0.055	0.055	0.055			
	(0.103)	(0.076)	(0.071)			
1(age > 55)*1995	-0.024	-0.024	-0.024			
	(0.101)	(0.080)	(0.076)			
1(age > 55)*1996	0.071	0.071	0.071			
	(0.102)	(0.077)	(0.075)			
1(age > 55)*1998	0.110	0.110	0.110			
	(0.092)	(0.077)	(0.071)			
1(age > 55)*1999	0.170	0.170*	0.170**			
	(0.120)	(0.092)	(0.082)			
1(age > 55)*2000	0.105	0.105	0.105*			
	(0.104)	(0.069)	(0.064)			
1(age > 55)*2001	0.197**	0.197***	0.197***			
	(0.098)	(0.071)	(0.065)			
1(age > 55)*2002	0.180*	0.180**	0.180***			
	(0.103)	(0.071)	(0.067)			
1(age > 55)*2003	0.173*	0.173**	0.173**			
	(0.104)	(0.075)	(0.073)			
metro area effects	N	Y				
year effects	Y	Υ				
metro-year effects	N	N	Y			
age category effects	Y	Y	Y			
R-squared	0.14	0.67	0.77			
N	1280	1280	1280			

Fig. 15: Coeffs. on age\*year interaction (Table 4 Col 3)



#### Conclusion

- Two basic points:
  - There is under-reporting. Third-party reporting does not eliminate evasion.
  - The extent of under-reporting appears to respond to economic incentives, in particular to change in employees' incentive to ensure accurate reporting.
- Implication: giving employees incentives to monitor employers should be a consideration in the design of social-insurance systems.
- Future work:
  - Estimating incidence of payroll taxes/social insurance benefits in presence of evasion.
  - Does greater compliance on intensive margin (less under-reporting by registered firms) induce lower compliance on extensive margin (fewer firms registering)?

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#### Related literature

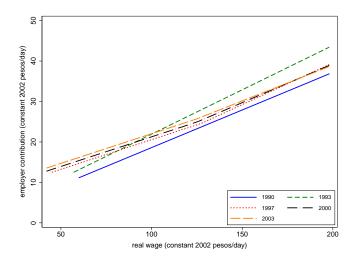
- Results broadly consistent with theoretical model in Kleven, Kreiner and Saez (2009).
  - Firms are cooperatives of workers who may collude in under-reporting.
  - In baseline case, all workers have access to firm's records, can credibly denounce evasion if it exists, cannot pre-commit not to denounce.
  - ▶ Two mechanisms make collusion more difficult in larger firms:
    - Workers subject to random shocks (e.g. becoming disgruntled).
    - Reward for whistle-blowing increasing in amount of evasion.
  - ► Tailorable to our setting:
    - Interpret pension reform as (small) increase in reward for whistle-blowing.
    - In presence of disgruntlement shocks, predicts (1) greater compliance in larger firms, (2) increased compliance by firms employing affected workers.

## Related literature (cont.)

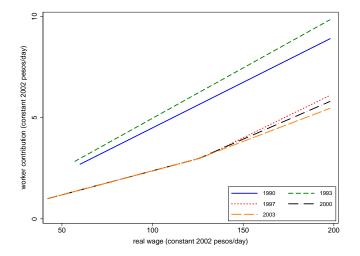
- Attractive enforcement properties of VATs:
  - ► Kopczuk and Slemrod (2006), Keen and Lockwood (2010), Pomeranz (2011).
- ► Papers using independent sources of information to infer under-reporting (not of payroll):
  - ► Fisman and Wei (2004), Gorodnichenko et al. (2009), Marion and Muehlegger (2008), Hurst et al. (2011), Braguinsky et al. (2010).
- Small literature on under-reporting of payroll:
  - ▶ Nyland et al. (2006), Tonin (2011), Elek et al. (forthcoming).
- Broader literatures on the role of firms in tax systems and on tax evasion/avoidance. Reviews:
  - ► Slemrod (2008), Gordon and Li (2009), Andreoni et al. (1998) Slemrod and Yitzhaki (2002), Saez et al. (2012).
- This paper appears to be the first empirical study of response of under-reporting by firms to changes in the incentives of employees.



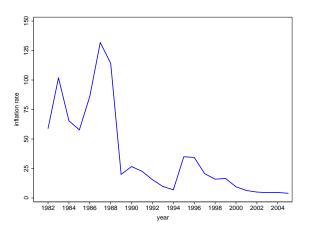
### Fig. 1: Employer contribution schedule (low wages)



### Fig. 2: Employee contribution schedule (low wages)



#### Inflation rate



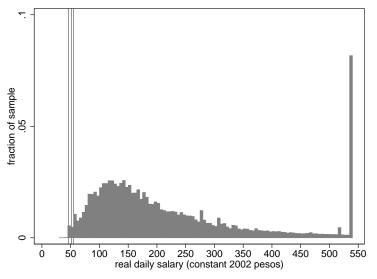
- ► Housing account
  - ► Employer contributes 5% of worker's wage to housing fund (INFONAVIT), to which workers can apply for loans.
  - Workers can claim unused funds at retirement.
    - ▶ Prior to 1992: *nominal* contributions, real value low.
    - ▶ 1992-1997: nominal contributions + interest, but real rate of return negative.
    - Post-reform: Funds administered by AFORE, can be claimed by workers who choose PRA.
    - Grandfathered workers who choose PAYGO only receive unused housing funds from 1992-1997.
  - Changes reinforce pension changes.

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    - Post-reform: Funds administered by AFORE, can be claimed by workers who choose PRA.
    - Grandfathered workers who choose PAYGO only receive unused housing funds from 1992-1997.
  - Changes reinforce pension changes.
- Enforcement:
  - Law provides for penalties as share of evasion (currently 40-100%), in addition to paying unpaid contributions.
  - No reward for whistle-blowers (beyond correcting employer contributions)
  - ► Wage must be corresponding minimum wage + 4.5% (to cover legally required annual bonus)



- Other dimensions of tax system:
  - VAT: 15% for 1988-2003 period.
  - Corporate income taxes:
    - ▶ 39.2% in 1988, 34% in 2003
    - Widspread evasion: e.g. in early 1990s, 70% of corporations declared no income (OECD, 1992).
  - Personal income taxes:
    - ▶ 3-50% in 1988, 3-34% in 2003.
    - Extensive tax credits for low-income workers, to offset regressive effects of VAT.
    - In 1997, individuals making <3.2 minimum wages (70% of all employees) paid ≤0 income tax (OECD, 1999, p. 80).</p>
  - VAT, social security taxes each ~3% of GDP; corporate + personal income taxes and PEMEX contributions each ~4% of GDP (OECD, 1999).
  - ► IMSS and tax authority first signed agreement to share data in June 2002. No information sharing previously.

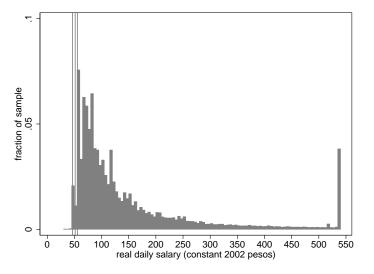
Fig. 9: Wage histogram, men, 1993, EIA plants



Notes: Bins are 2 pesos wide. Average 2002 exchange rate: 9.66 pesos/dollar.



Fig. 10: Wage histogram, men, 1993, EMIME plants





## Table 4: Differential effects on employment gap, men

	(1)	(2)	(3)
1(age > 55)*1988	-0.026	-0.026	-0.026
,	(0.105)	(0.098)	(0.100)
1(age > 55)*1989	0.048	0.048	0.048
,	(0.103)	(0.097)	(0.103)
1(age > 55)*1990	0.077	0.077	0.077
	(0.096)	(0.088)	(0.097)
1(age > 55)*1991	0.109	0.109	0.109
	(0.111)	(0.110)	(0.109)
1(age > 55)*1992	0.054	0.054	0.054
	(0.101)	(0.096)	(0.100)
1(age > 55)*1993	0.098	0.098	0.098
	(0.092)	(0.087)	(0.091)
1(age > 55)*1994	-0.224**	-0.224**	-0.224**
	(0.098)	(0.095)	(0.101)
1(age > 55)*1995	0.029	0.029	0.029
	(0.112)	(0.105)	(0.107)
1(age > 55)*1996	0.005	0.005	0.005
	(0.102)	(0.100)	(0.102)
1(age > 55)*1998	0.045	0.045	0.045
	(0.106)	(0.099)	(0.104)
1(age > 55)*1999	0.031	0.031	0.031
	(0.112)	(0.104)	(0.106)
1(age > 55)*2000	-0.006	-0.006	-0.006
	(0.094)	(0.093)	(0.093)
1(age > 55)*2001	0.014	0.014	0.014
	(0.110)	(0.109)	(0.105)
1(age > 55)*2002	0.091	0.091	0.091
	(0.113)	(0.107)	(0.103)
1(age > 55)*2003	0.034	0.034	0.034
	(0.094)	(0.093)	(0.091)
metro area effects	N	Y	
year effects	Y	Y	
metro-year effects	N	N	Y
age category effects	Y	Y	Υ
R-squared	0.28	0.37	0.54
N	1280	1280	1280

Did reform affect wage difference between formal and informal sectors?

- Did reform affect wage difference between formal and informal sectors?
- Estimate a Mincer-type wage equation, separately by age group and year:

$$\ln w_i = \alpha + \beta (has\_imss_i) + X_i \gamma + \varepsilon_i$$

where  $X_i$  includes age, indicators for married, 9 schooling categories, 22 occupational categories, 16 metropolitan areas, 50 sectors

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- Estimate a Mincer-type wage equation, separately by age group and year:

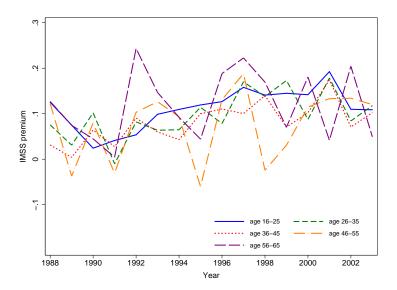
$$\ln w_i = \alpha + \beta (has\_imss_i) + X_i \gamma + \varepsilon_i$$

where  $X_i$  includes age, indicators for married, 9 schooling categories, 22 occupational categories, 16 metropolitan areas, 50 sectors

▶ Call  $\widehat{\beta}$  the "IMSS premium."

- ▶ All else equal, we would expect  $\beta$  < 0: covered workers pay for benefits in form of lower wage.
- ▶ If workers are aware of under-reporting, we would expect  $\beta \downarrow$  relatively more for younger workers with reform, since pension value of being covered  $\uparrow$ .

#### IMSS "premium", by age group over time



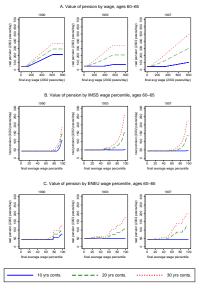
#### Discussion: IMSS "premium"

- $ightharpoonup \widehat{\beta}$ 's generally positive, increasing over period.
  - Suggests positive selection on unobservables into formal-sector jobs.
  - N.B.: Sample is same as above, does not include self-employed, may understate payoff in informal sector. (May explain difference with Marcouiller et al. (1997).)

#### Discussion: IMSS "premium"

- $ightharpoonup \widehat{\beta}$ 's generally positive, increasing over period.
  - Suggests positive selection on unobservables into formal-sector jobs.
  - ▶ N.B.: Sample is same as above, does not include self-employed, may understate payoff in informal sector. (May explain difference with Marcouiller et al. (1997).)
- ▶ No evidence of differential effect on IMSS "premium" by age.
  - Suggestive that workers were not aware of under-reporting, became more willing to pay (now reduced) cost of figuring out extent of under-reporting.
  - Need research design with exogenous variation in formality/informality at worker level to answer definitively.

#### Fig. A1: Value of pension, women ages 60-65



#### Table A5: Comparison of IMSS and ENEU, 1990, women

A. 1990	IMSS baseline sample (1)	full ENEU sample (2)	ENEU w/ IMSS (3)	ENEU w/o IMSS (4)	ENEU permanent w/ IMSS (5)	ENEU full-time w/ IMSS (6)
	11400	100.66	107.00	104.04		100.00
real avg daily wage	114.29	133.66	137.03	124.94		128.68
	(0.10)	(2.16)	(2.65)	(3.59)		(2.51)
age	28.16	28.35	28.03	29.17		27.82
	(0.01)	(0.21)	(0.23)	(0.47)		(0.24)
fraction employed in ests >100 employees	0.54	0.45	0.54	0.21		0.54
	(0.00)	(0.01)	(0.01)	(0.02)		(0.01)
N (raw observations)	815760	6685	5126	1559		4745
N (population, using weights)	815760	1023858	738698	285160		677053
B. 2000						
real avg daily wage	119.01	128.15	135.99	109.81	140.67	129.75
	(0.09)	(1.82)	(2.22)	(3.06)	(2.49)	(2.19)
age	30.50	30.34	29.85	31.50	30.17	29.71
-	(0.01)	(0.18)	(0.19)	(0.40)	(0.21)	(0.20)
fraction employed in ests >100 employees	0.63	0.49	0.62	0.19	0.64	0.62
	(0.00)	(0.01)	(0.01)	(0.01)	(0.01)	(0.01)
N (raw observations)	1267196	9670	7227	2443	6305	6607
N (population, using weights)	1267196	1652164	1157184	494980	1001866	1056013
(Population, asing Weights)	120,130	1002101	1101101	.5.500	1001000	1000010

→ Return

#### Fig. A2: Employment, IMSS vs. ENEU samples, women

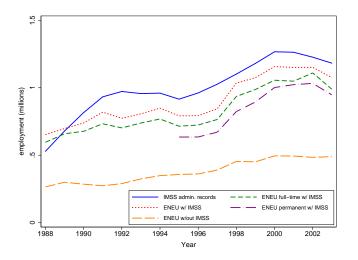




Fig. A3: Wage histograms, women, 1990

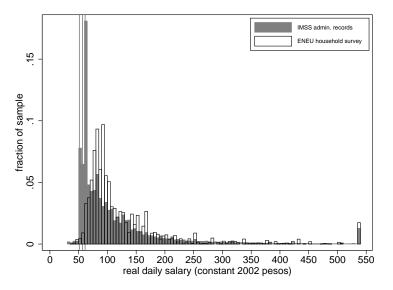
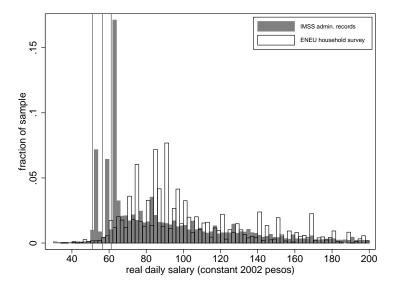




Fig. A4: Wage histograms, women, 1990, low wages





#### Fig. A5: Wage histograms, women, 1990, by firm size

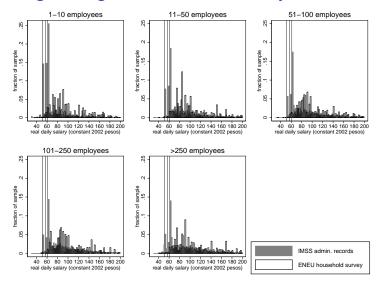
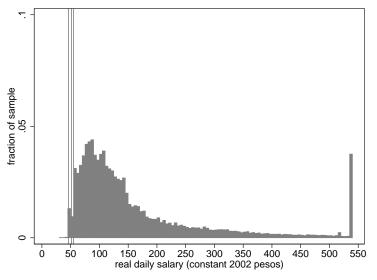




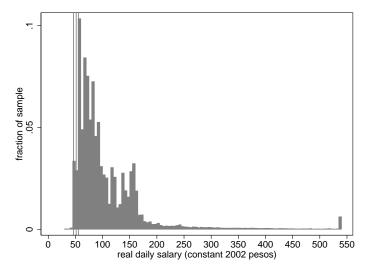
Fig. A6: Wage histogram, women, 1993, EIA plants





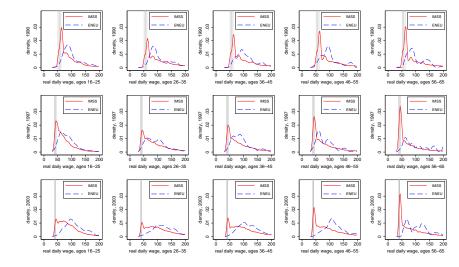


#### Fig. A7: Wage histogram, women, 1993, EMIME plants





#### Fig. A8: Wage densities by age group, women



#### Fig. A9: Wage gaps by age group, women

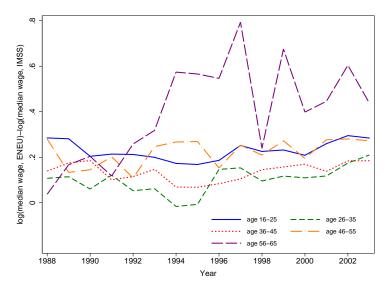




Fig. A10: Wage gaps by age group, women, deviated from metro-year means

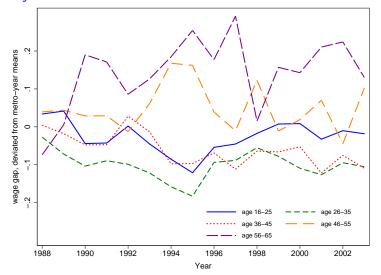


Fig. A11: Kullback-Liebler divergence by age group, women

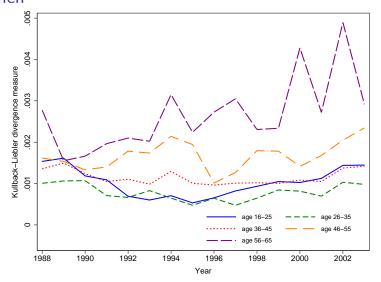
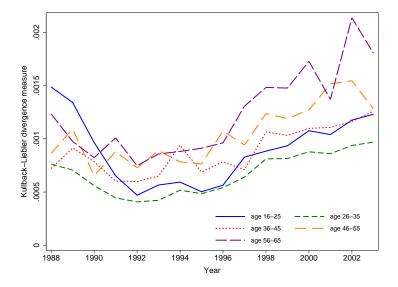


Fig. 14: Kullback-Liebler divergence by age group, men

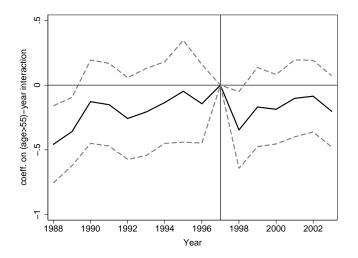




### Table A6: Differential effects on wage gap, women dep. var.: log(median wage, ENEU) - log(median wage, IMSS)

	(1)	(2)	(3)
1(age > 55)*1988	-0.477***	-0.457***	-0.457***
	(0.178)	(0.164)	(0.152)
1(age > 55)*1989	-0.362**	-0.370**	-0.358***
	(0.158)	(0.155)	(0.134)
1(age > 55)*1990	-0.147	-0.123	-0.127
	(0.191)	(0.177)	(0.164)
1(age > 55)*1991	-0.167	-0.159	-0.151
	(0.207)	(0.188)	(0.163)
1(age > 55)*1992	-0.283	-0.267	-0.257
	(0.185)	(0.180)	(0.161)
1(age > 55)*1993	-0.219	-0.211	-0.207
	(0.198)	(0.189)	(0.172)
1(age > 55)*1994	-0.180	-0.167	-0.134
	(0.182)	(0.178)	(0.161)
1(age > 55)*1995	-0.066	-0.060	-0.047
	(0.216)	(0.218)	(0.201)
1(age > 55)*1996	-0.155	-0.149	-0.143
	(0.186)	(0.175)	(0.155)
1(age > 55)*1998	-0.363**	-0.350**	-0.346**
	(0.179)	(0.165)	(0.152)
1(age > 55)*1999	-0.185	-0.177	-0.169
	(0.185)	(0.174)	(0.156)
1(age > 55)*2000	-0.197	-0.185	-0.186
	(0.176)	(0.159)	(0.137)
1(age > 55)*2001	-0.114	-0.108	-0.102
	(0.186)	(0.174)	(0.152)
1(age > 55)*2002	-0.097	-0.091	-0.085
	(0.173)	(0.161)	(0.141)
1(age > 55)*2003	-0.214	-0.208	-0.202
	(0.163)	(0.156)	(0.140)
metro area effects	N	Y	
year effects	Y	Y	
metro-year effects	N	N	Υ
age category effects	Υ	Y	Y
R-squared	0.14	0.34	0.50
N	1258	1258	1258

Fig. A12: Coeffs. on age\*year interaction (Table 4 Col 3)

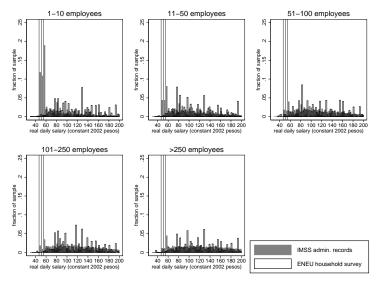




### Table A7: Differential effects on employment gap, women

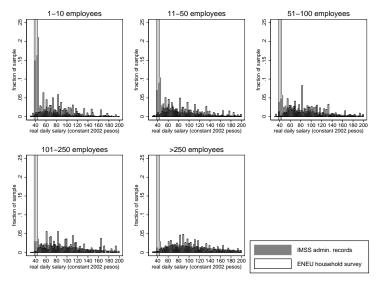
	dep. var.:	dep. var.: log(employment, IMSS) - log(employment, ENEU)				
	(1)	(2)	(3)			
1(age > 55)*1988	-0.141	-0.176	-0.185			
,	(0.252)	(0.237)	(0.229)			
1(age > 55)*1989	0.161	0.153	0.186			
,	(0.234)	(0.222)	(0.194)			
1(age > 55)*1990	0.139	0.129	0.153			
,	(0.238)	(0.219)	(0.199)			
1(age > 55)*1991	0.246	0.243	0.244			
,	(0.220)	(0.215)	(0.201)			
1(age > 55)*1992	-0.172	-0.174	-0.174			
, - ,	(0.265)	(0.259)	(0.236)			
1(age > 55)*1993	0.156	0.169	0.165			
,	(0.234)	(0.230)	(0.222)			
1(age > 55)*1994	0.029	0.019	-0.014			
,	(0.260)	(0.244)	(0.232)			
1(age > 55)*1995	-0.331	-0.321	-0.314			
, - ,	(0.285)	(0.271)	(0.255)			
1(age > 55)*1996	-0.095	-0.093	-0.091			
, - ,	(0.240)	(0.222)	(0.207)			
1(age > 55)*1998	-0.133	-0.115	-0.116			
	(0.203)	(0.191)	(0.183)			
1(age > 55)*1999	-0.286	-0.295	-0.289			
	(0.255)	(0.239)	(0.220)			
1(age > 55)*2000	-0.153	-0.163	-0.153			
	(0.257)	(0.238)	(0.221)			
1(age > 55)*2001	0.144	0.146	0.148			
	(0.225)	(0.211)	(0.201)			
1(age > 55)*2002	-0.013	-0.011	-0.009			
	(0.300)	(0.286)	(0.260)			
1(age > 55)*2003	-0.275	-0.272	-0.271			
	(0.245)	(0.245)	(0.223)			
metro area effects	N	Y				
year effects	Υ	Y				
metro-year effects	N	N	Y			
age category effects	Υ	Υ	Y			
R-squared	0.23	0.32	0.46			
N	1258	1258	1258			

#### Wage histograms, men, 1993, by firm size



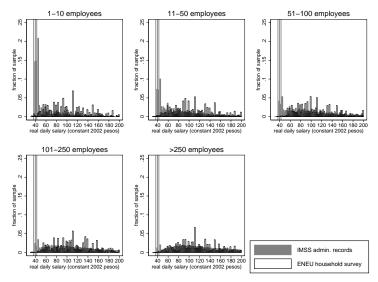


#### Wage histograms, men, 1997, by firm size



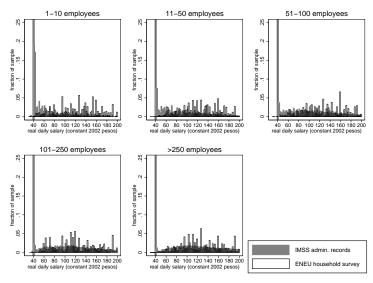


#### Wage histograms, men, 2000, by firm size



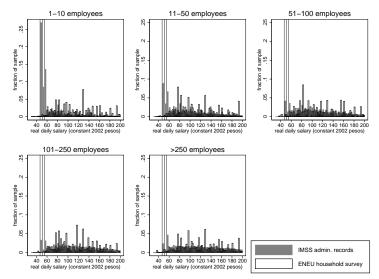


#### Wage histograms, men, 2003, by firm size



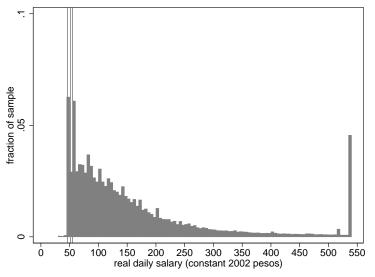


#### Wage histograms, men, 1993, by firm size, non-EIA plants



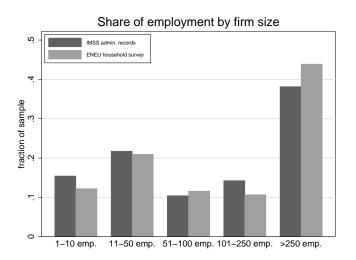


#### Wage histogram, men, 1993, non-EIA plants

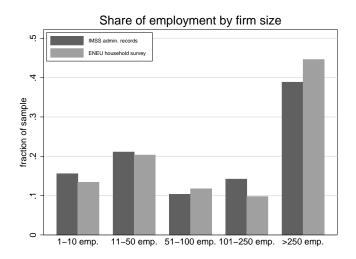


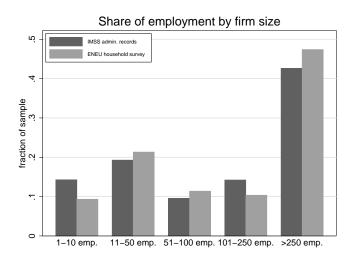


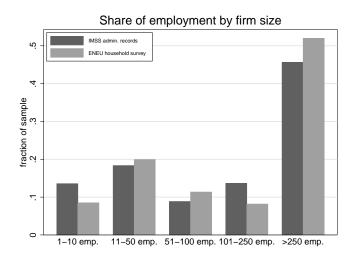


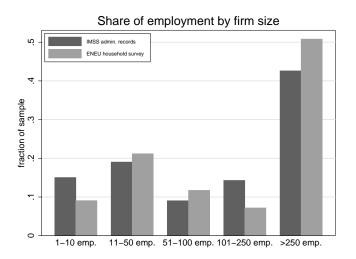






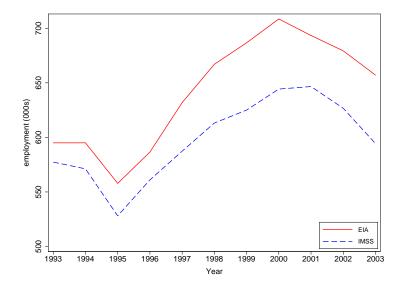




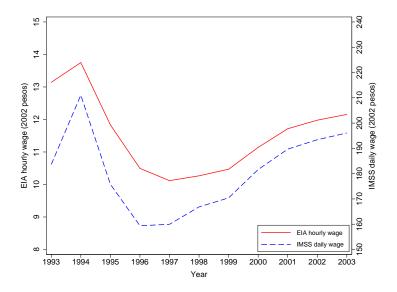




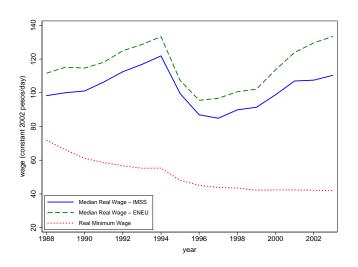
#### Employment, IMSS vs. EIA



#### Wages, IMSS vs. EIA

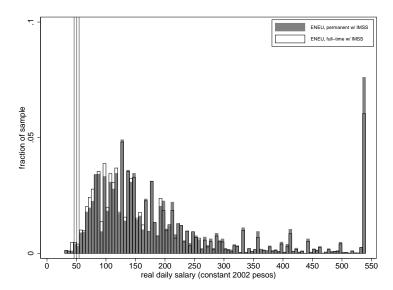


#### Mean, median, minimum wages

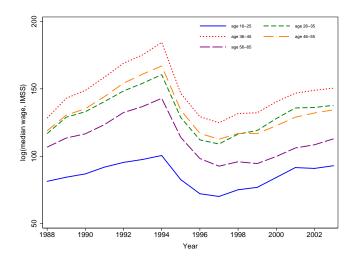




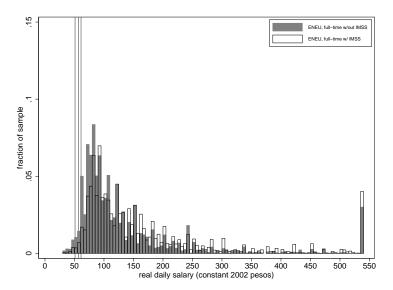
# ENEU wage distributions, full-time vs. permanent w/IMSS, men, 1994



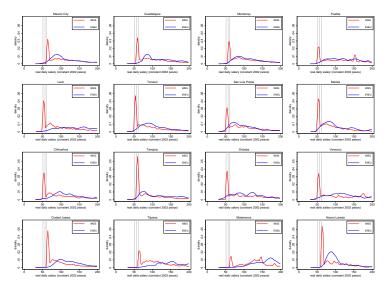
#### Log median daily wages, men, IMSS data, by age group



## Wage histograms, covered vs. not covered by IMSS, men, 1990



#### Wage distributions, by metro area, men, 1990



#### Wage gaps (in means) by age group, men

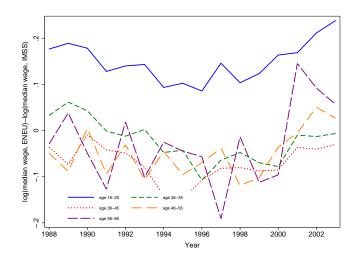


Table 1: Tenure in IMSS system, 1997 Q2, baseline sample

– Years	Men			Women						
	16-25	26-35	36-45	46-55	56-65	16-25	26-35	36-45	46-55	56-65
in IMSS	(%)	(%)	(%)	(%)	(%)	(%)	(%)	(%)	(%)	(%)
0	27.9	6.7	4.4	4.4	6.1	29.6	10.0	8.0	5.9	6.3
1	23.0	8.0	4.6	4.4	5.8	24.0	11.2	8.4	5.8	6.1
2	14.1	7.4	4.1	3.7	4.4	14.4	9.4	6.8	4.7	4.4
3	11.7	8.0	4.4	3.7	4.1	11.5	9.5	7.1	5.3	5.5
4	8.9	8.3	4.6	3.9	4.3	8.3	9.2	6.9	5.3	5.3
5	6.7	9.1	5.2	4.3	4.5	5.9	9.4	7.1	5.6	5.1
6	4.5	10.5	7.3	6.3	6.1	3.7	9.8	8.3	7.8	7.6
7	2.3	9.4	6.4	5.5	5.2	1.8	8.6	7.0	6.8	6.1
8	0.8	8.6	6.5	5.4	4.9	0.7	7.1	6.4	6.4	5.9
9	0.1	7.3	9.0	9.7	10.1	0.1	5.4	6.9	8.1	8.8
10	0.0	5.6	7.4	6.3	4.8	0.0	3.7	5.4	5.5	4.3
11	0.0	5.2	9.8	8.7	6.8	0.0	3.2	6.2	7.0	5.7
12	0.0	5.9	26.5	33.5	32.9	0.0	3.5	15.7	25.8	29.2
N (000s)	646.3	767.3	412.3	198.2	71.8	425.1	355.6	165.0	63.1	17.9